

Stationary and Wide Sense Stationary Processes

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Definition 1. A RP $\mathcal{X} = \{X_t : t \in \mathbb{R}\}$ or $\mathcal{X} = \{X_n : n = \dots, -2, -1, 0, 1, 2, \dots\}$ is stationary if all its k -order marginals do not depend on a translation by a constant: for any $k > 0$ and for any time indices t_1, \dots, t_k ,

$$\forall \tau \in \mathbb{R} \quad f_{X_{t_1}, \dots, X_{t_k}}(x_1, \dots, x_k) = f_{X_{t_1+\tau}, \dots, X_{t_k+\tau}}(x_1, \dots, x_k).$$

The above expression is for continuous-valued RP. If the RP is discrete valued, replace the pdfs with pmfs.

Definition 2. A RP $\mathcal{X} = \{X_t : t \in \mathbb{R}\}$ or $\mathcal{X} = \{X_n : n = \dots, -2, -1, 0, 1, 2, \dots\}$ is wide sense stationary (WSS) if its mean function $m_{\mathcal{X}}(t)$ is constant and its autocorrelation $R_{\mathcal{X}}(t, s)$ is a function of $|s - t|$. In this case, we abuse notation (by wide convention) and consider the autocorrelation as a function of one variable - the difference between the two time points $R_{\mathcal{X}}(\tau) = R_{\mathcal{X}}(t, t + \tau)$.

A stationary RP is also WSS but the converse is not true in general. WSS processes will prove to be useful for signal processing purposes as they simplify some expressions in a considerable manner. Note that the autocovariance of a WSS RP is also a function of the difference between the two time points.

Below are some properties of the autocorrelation function of WSS RP \mathcal{X} :

1. $R_{\mathcal{X}}(0) = \mathbf{E}(X_t^2)$ for all t (it measures the average “power” of the process which is fixed for all t). Since X_t^2 is a non-negative RV, $R_{\mathcal{X}}(0) \geq 0$.

2. $R_{\mathcal{X}}$ is an even function

$$R_{\mathcal{X}}(\tau) = \mathbf{E}(X_t X_{t+\tau}) = \mathbf{E}(X_{t+\tau} X_t) = R_{\mathcal{X}}(-\tau).$$

3. The probability of a change in the RP realization after some time is bounded by the difference in the autocorrelation function:

$$\begin{aligned} P(|X_{t+\tau} - X_t| \geq \epsilon) &= P((X_{t+\tau} - X_t)^2 \geq \epsilon^2) \leq \frac{\mathbf{E}((X_{t+\tau} - X_t)^2)}{\epsilon^2} \\ &= \frac{\mathbf{E}(X_{t+\tau}^2 + X_t^2 - 2X_{t+\tau}X_t)}{\epsilon^2} = \frac{2(R_{\mathcal{X}}(0) - R_{\mathcal{X}}(\tau))}{\epsilon^2} \end{aligned}$$

where the inequality follows from Markov inequality.

4. The autocorrelation function attains its maximum at $\tau = 0$.

First, we show that since the expectation of a non-negative RV is non-negative:

$$\begin{aligned} 0 &\leq \mathbf{E}((X/\sqrt{\mathbf{E}(X^2)} - Y/\sqrt{\mathbf{E}(Y^2)})^2) = \mathbf{E}(X^2/\mathbf{E}(X^2) + Y^2/\mathbf{E}(Y^2) - 2XY/\sqrt{\mathbf{E}(X^2)\mathbf{E}(Y^2)}) \\ &= 1 + 1 - 2\mathbf{E}(XY)/\sqrt{\mathbf{E}(X^2)\mathbf{E}(Y^2)} = 2(1 - \mathbf{E}(XY)/\sqrt{\mathbf{E}(X^2)\mathbf{E}(Y^2)}) \Rightarrow \mathbf{E}(XY) \leq \sqrt{\mathbf{E}(X^2)\mathbf{E}(Y^2)} \end{aligned}$$

and hence $(\mathbf{E}(XY))^2 \leq \mathbf{E}(X^2)\mathbf{E}(Y^2)$. Using this last result we have

$$(R_{\mathcal{X}}(\tau))^2 = (\mathbf{E}(X_{t+\tau}X_t))^2 \leq \mathbf{E}((X_{t+\tau})^2)\mathbf{E}((X_t)^2) = (R_{\mathcal{X}}(0))^2.$$

The above result, together with the fact that $R_{\mathcal{X}}(0) = \mathbf{E}(X_t^2)$ is non-negative (property 1) yields the required result: $R_{\mathcal{X}}(0) \geq R_{\mathcal{X}}(\tau)$ for all τ .